

Q2a) We need to show  $E\{ |X_t| \} < \infty$  for all  $t$ , (i)

and  $E(X_{t+1} | X_t, X_{t-1}, \dots, X_0) = X_t$ . (ii)

i) Clearly,  $0 \leq X_t \leq N$  for all  $t$ , so  $E\{ |X_t| \} < \infty$ .

ii)  $[X_{t+1} | X_t] \sim \text{Binomial}(N, \frac{X_t}{N})$

so  $E(X_{t+1} | X_t, \dots, X_0) = E(X_{t+1} | X_t)$  (Markov Property)  
 $= N * \frac{X_t}{N}$  (Binomial mean)  
 $= X_t$  as required.

Thus  $X_0, X_1, \dots$  is a martingale.

b)  $T = \min \{ t : X_t = 0 \text{ or } X_t = N \}$ .

So  $T$  depends only on events visible up to and including time  $T$  not on any future events.

This is the requirement for  $T$  to be a stopping time w.r.t.  $\{X_t\}$ .

c) For the Optional Stopping Theorem, we need:

i)  $E T < \infty$  (given in Q)

ii) There exists some constant  $c$  such that

$E(|X_{t+1} - X_t|) \leq c$  for all  $t$ .

This is satisfied putting  $c=N$ , because

$0 \leq X_t \leq N$  for all  $t$ .

So we can apply the Optional Stopping Theorem.

2c cont.) By O.S.T. and because  $T$  is a stopping time w.r.t.  $\{X_t\}$ , (2)

$E(X_T) = E(X_0) = x$  in this case. (2)

Now let  $p = P(\text{eventually fixed at state } N)$ .

By definition of  $T$ ,

$X_T = \begin{cases} N & \text{with prob. } p \\ 0 & \text{w.p. } 1-p \end{cases}$

So  $E(X_T) = pN + (1-p)*0 = pN$ . (1)

Equating (1) and (2) gives  $x = pN$

$\Rightarrow p = \frac{x}{N}$ .

This is the probability the population is eventually fixed for allele  $A$ , i.e. chain stops in state  $N$ .

Q4a) First-step analysis:  $h_x = \frac{2}{3}h_{x+1} + \frac{1}{3}h_{x+1}$   $x=1, \dots, N-1$

So  $h_{x+1} - 3h_x + 2h_{x-1} = 0$  (\*)

Valid for  $x=1, \dots, N-1$ .

Boundaries:  $h_0 = 0$   $h_N = 1$ .

The eqn (\*) is homogeneous, so we can proceed.

Characteristic equation:  $t^2 - 3t + 2 = 0$

$(t-1)(t-2) = 0$

$\Rightarrow t=1, t=2$  distinct real roots.

4a cont.) Deduce solutions to  $\textcircled{*}$  of form  $\begin{cases} h_x = 1^x = 1 \\ h_x = 2^x \end{cases}$

Check for linear independence:

$$\det \begin{pmatrix} 1^0 & 2^0 \\ 1^1 & 2^1 \end{pmatrix} = 2 - 1 = 1 \neq 0.$$

Thus  $\textcircled{*}$  has two linearly independent solutions which form a basis for the 2-dimensional solution space.

$\therefore$  General solution of  $\textcircled{*}$ :

$$h_x = A + B 2^x \text{ for } x=0, 1, 2, \dots, N. \quad \textcircled{**}$$

$A, B$  constants to be found.

Boundary conditions:

$$h_0 = A + B = 0 \quad \textcircled{1}$$

$$h_N = A + B 2^N = 1 \quad \textcircled{2}$$

$$\textcircled{2} - \textcircled{1}: \quad B(2^N - 1) = 1$$

$$\therefore B = \frac{1}{2^N - 1}, \quad A = -B.$$

Final solution to  $\textcircled{*}$ :

$$h_x = B(2^x - 1)$$

$$\therefore h_x = \frac{2^x - 1}{2^N - 1} \text{ for } x=0, 1, \dots, N$$

as stated.

(3)

4b) First-step analysis:

$$m_x = 1 + \frac{1}{3} m_{x+1} + \frac{2}{3} m_{x-1} \quad x=1, \dots, N-1$$

$$\Rightarrow m_{x+1} - 3m_x + 2m_{x-1} = -3 \quad \textcircled{E}$$

for  $x=1, \dots, N-1$ .

Boundaries:  $m_0 = m_N = 0$ . (given).

This is an inhomogeneous eqn, but we have already solved the corresponding homogeneous eqn in (a).

Homogeneous equation:

$$\textcircled{H} \quad u_{x+1} - 3u_x + 2u_{x-1} = 0 \text{ same as } \textcircled{*} \text{ part (a)}$$

General solution to  $\textcircled{H}$ :

$$u_x = C + D 2^x \quad x=0, \dots, N.$$

from  $\textcircled{**}$  in (a).

$C, D$  constants to be found.

Particular solution to  $\textcircled{E}$ :

Try  $m_x = \theta x$ :

$$\Rightarrow \theta(x+1) - 3\theta x + 2\theta(x-1) = -3$$

$$\theta(x-3x+2x) + \theta(1-2) = -3$$

$$\Rightarrow \theta = 3.$$

So  $m_x = 3x$  is a particular solution to  $\textcircled{E}$ .

(4)

4b cont.) General soln to (E):

$$M_x = C + D 2^x + 3x \quad x=0, \dots, N.$$

Boundaries:  $M_0 = C + D = 0$  (1)

$$M_N = C + D 2^N + 3N = 0 \quad (2)$$

$$(2) - (1) \Rightarrow D(2^N - 1) = -3N$$

$$\therefore D = \frac{-3N}{2^N - 1}, \quad C = -D.$$

Final solution to (E):

$$M_x = \frac{-3N}{2^N - 1} (2^x - 1) + 3x$$

$$\Rightarrow M_x = 3x - 3N \frac{(2^x - 1)}{2^N - 1} \quad x=0, \dots, N.$$

4c) For the event  $\{Y \geq y\}$ , consider the new transition diagram:



This is the same system as in (a) but with  $y$  replacing  $N$ .

So  $P(Y \geq y) = P(\text{chain hits } y \text{ before hitting } 0)$

$$= \frac{2^x - 1}{2^y - 1} \quad \text{as long as } y \geq x.$$

For the event  $\{Y = y\}$  we need the chain to hit  $y$  and not to hit  $y+1$ :  $P(Y=y) = P(Y \geq y) - P(Y \geq y+1)$  for  $y=x, \dots, N-1$ .

(5)

4c cont.) Overall:

$$P(Y=y) =$$

$$\begin{cases} 0 & y=0, \dots, x-1 \\ \frac{2^x - 1}{2^y - 1} - \frac{2^x - 1}{2^{y+1} - 1} & y=x, \dots, N-1 \\ \frac{2^x - 1}{2^N - 1} & y=N. \end{cases}$$

Starting from state  $x$ .

(6)