

(1)

1a) A and B are independent  $\Leftrightarrow P(A \cap B) = P(A)P(B)$ .

$$\begin{aligned} \text{Now } P(A \cap \bar{B}) &= P(A) - P(A \cap B) \text{ by Partition Rule} \\ &= P(A) - P(A)P(B) \text{ by independence} \\ &= P(A)(1 - P(B)) \\ &= P(A)P(\bar{B}). \end{aligned}$$

Thus A and  $\bar{B}$  are independent, by definition.

b) Consider  $A \cap B$ : event  $(A \cap B)$  is a subset of A and a subset of B, so

$$\begin{aligned} P(A \cap B) &\leq P(A) = \frac{4}{5} \\ &\text{and} \\ P(A \cap B) &\leq P(B) = \frac{1}{2}. \end{aligned}$$

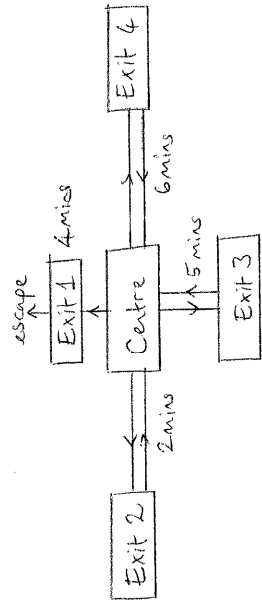
Only the smaller inequality is necessary, so  $P(A \cap B) \leq \frac{1}{2}$ .

Now consider  $A \cup B$ :

$$\begin{aligned} P(A \cup B) &= P(A) + P(B) - P(A \cap B) \leq 1 \\ \Rightarrow \frac{4}{5} + \frac{1}{2} - P(A \cap B) &\leq 1 \\ \frac{4}{5} + \frac{1}{2} - 1 &\leq P(A \cap B) \\ \Rightarrow \frac{3}{10} &\leq P(A \cap B). \end{aligned}$$

Thus overall we have  $\frac{3}{10} \leq P(A \cap B) \leq \frac{1}{2}$  as required.

(2)



2)

a) Using ideas of conditional expectation:

$$\begin{aligned} \mathbb{E}(T) &= \frac{1}{4} \{ \mathbb{E}(T | \text{exit 1}) + \dots + \mathbb{E}(T | \text{exit 4}) \} \\ &= \frac{1}{4} \{ 4 + (2 + \mathbb{E}T) + (5 + \mathbb{E}T) + (6 + \mathbb{E}T) \} \\ &= \frac{1}{4} \cdot 17 + \frac{3}{4} \mathbb{E}(T) \\ \frac{1}{4} \mathbb{E}(T) &= \frac{17}{4} \end{aligned}$$

$$\Rightarrow \underline{\underline{\mathbb{E}(T) = 17 \text{ mins.}}}$$

b) There are 3 exits that take time  $\geq 3$  minutes, all equally likely. Thus  $P(\text{Exit 1} | \text{time} \geq 3 \text{ mins}) = \frac{1}{3}$ .

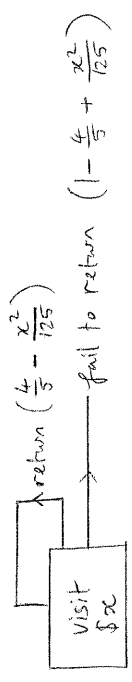
$$c) P(\text{escapes in 6 minutes}) = P(\text{chooses Exit 2 first, then chooses Exit 1})$$

because this is the only combination of exits that allow escape in 6 minutes.

$$\text{Thus } \underline{\underline{P(\text{escapes in 6 mins}) = \frac{1}{4} * \frac{1}{4} = \frac{1}{16}.}}$$

③

3) a) Profit per visit =  $\$x$ .



$$EP = x + \left(\frac{x}{5} - \frac{x^2}{125}\right) EP$$

$$\Rightarrow EP \left\{ 1 - \frac{x}{5} + \frac{x^2}{125} \right\} = x$$

$$EP = \frac{x}{\frac{1}{5} + \frac{x^2}{125}}$$

$$EP = \frac{125x}{25+x^2}; \text{ as stated.}$$

b) Let  $g(x) = \frac{125x}{25+x^2} = 125x(25+x^2)^{-1}$

Then the optimal  $x$  occurs where  $g'(x) = 0$ .

$$g'(x) = 125(25+x^2)^{-1} + 125x(-1)(25+x^2)^{-2}(2x) = 0$$

$$\Rightarrow \frac{125}{(25+x^2)^2} \{ 25+x^2 - 2x^2 \} = 0$$

$$\Rightarrow 25+x^2 - 2x^2 = 0$$

$$x^2 = 25$$

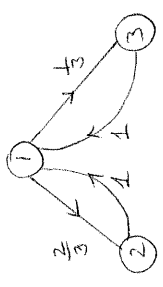
$$x = 5 \quad (\text{need } x > 0)$$

Question states that any stationary point in  $0 \leq x \leq 10$  is a maximum.

Thus the optimal profit per visit is  $\underline{\underline{\$5}}$ .

④

4a)



b) Equilibrium distribution,  $\pi$ , satisfies  $\pi^T P = \pi^T$  and  $\pi_1 + \pi_2 + \pi_3 = 1$ .

$$\text{Thus } (\pi_1 \ \pi_2 \ \pi_3) \begin{pmatrix} 0 & \frac{2}{3} & \frac{1}{3} \\ \frac{1}{3} & 1 & 0 \\ 1 & 0 & 0 \end{pmatrix} = (\pi_1 \ \pi_2 \ \pi_3)$$

$$\Rightarrow \pi_2 + \pi_3 = \pi_1 \quad \text{①}$$

$$\frac{2}{3}\pi_1 = \pi_2 \quad \text{②}$$

$$\frac{1}{3}\pi_1 = \pi_3 \quad \text{③}$$

Substituting for  $\pi_2, \pi_3$  in eqn.  $\pi_1 + \pi_2 + \pi_3 = 1$

$$\Rightarrow \pi_1 + \frac{2}{3}\pi_1 + \frac{1}{3}\pi_1 = 1$$

$$\Rightarrow \pi_1 = \frac{1}{2}$$

$$\text{Thus } \underline{\underline{\pi = \left(\frac{1}{2}, \frac{1}{3}, \frac{1}{6}\right)}}$$

c) The matrix  $P$  is irreducible, but all states have period 2 so it is not aperiodic.

Thus  $X_t$  does not converge to  $\pi$  (or any other distribution) as  $t \rightarrow \infty$ .

5

$$\begin{aligned}
 5) a) \quad G_T(s) &= \mathbb{E}(s^T) \\
 &= \mathbb{E}(s^{X_1 + X_2 + \dots + X_N}) \\
 &= \mathbb{E}_N \{ \mathbb{E}(s^{X_1 + \dots + X_N} | N) \} \\
 &= \mathbb{E}_N \{ \mathbb{E}(s^{X_1} s^{X_2} \dots s^{X_N} | N) \} \\
 &= \mathbb{E}_N \{ \mathbb{E}(s^{X_1}) \mathbb{E}(s^{X_2}) \dots \mathbb{E}(s^{X_N}) \} \\
 &\quad \text{because the } X_i \text{'s are independent of each other and of } N \\
 &= \mathbb{E}_N \{ \mathbb{E}(s^{X_1})^N \} \quad \text{because the } X_i \text{'s are identically distributed} \\
 &= \mathbb{E}_N \{ G_X(s)^N \} \\
 \underline{G_T(s)} &= \underline{G_N(G_X(s))} \quad \text{by definitions.}
 \end{aligned}$$

b)  $Z_n$  is the sum of offspring of the  $Z_{n-1}$  individuals at time  $n-1$ . These offspring can be labelled  $Y_1, Y_2, \dots, Y_{Z_{n-1}}$ .

$$\text{Thus } Z_n = Y_1 + Y_2 + \dots + Y_{Z_{n-1}} \quad \left( \begin{array}{l} \text{sum of a random number} \\ \text{of random variables.} \end{array} \right)$$

$$\begin{aligned}
 \text{Put } T = Z_n &\Rightarrow G_T = G_N \\
 N = Z_{n-1} &\Rightarrow G_N = G_{n-1} \\
 X_i = Y_i &\Rightarrow G_X = G
 \end{aligned}$$

} in part (a).

$$\text{Then } G_T(s) = G_N(G_X(s))$$

$$\Rightarrow \underline{G_n(s) = G_{n-1}(G(s))} \quad \text{as required.}$$

c)  $P(Z_n = 1) = G'_n(0)$  (general property of PGF: see Attachment).

$$\text{Now } G'_n(s) = G'_{n-1}(G(s)) \quad G'(s)$$

$$\begin{aligned}
 \text{and } G(0) &= P(Y=0) = p_0 \\
 G'(0) &= P(Y=1) = p_1
 \end{aligned}$$

6

5c cont.) Thus

$$\begin{aligned}
 P(Z_n = 1) &= G'_n(0) \\
 &= G'_{n-1}(G(0)) \quad G'(0) \\
 &= G'_{n-1}(p_0) \quad p_1 \\
 \underline{P(Z_n = 1)} &= \underline{p_1 G'_{n-1}(p_0)} \quad \text{as required.}
 \end{aligned}$$

$$\text{Similarly, } P(Z_n = 2) = \frac{1}{2!} G''_n(0)$$

$$\text{But } G''_n(s) = G''_{n-1}(G(s)) \{G'(s)\}^2 + G'_{n-1}(G(s)) \cdot G''(s)$$

$$\text{and } G(0) = p_0, \quad G'(0) = p_1, \quad G''(0) = 2! p_2 = 2p_2.$$

$$\text{So } P(Z_n = 2) = \frac{1}{2!} \{ G''_{n-1}(G(0)) \{G'(0)\}^2 + G'_{n-1}(G(0)) \cdot G''(0) \}$$

$$= \frac{1}{2} \{ G''_{n-1}(p_0) p_1^2 + G'_{n-1}(p_0) \cdot 2p_2 \}$$

$$\underline{P(Z_n = 2) = \frac{1}{2} p_1^2 G''_{n-1}(p_0) + p_2 G'_{n-1}(p_0)}.$$

7

6) a)  $G(s) = E(s^Y)$

$$= \sum_{y=0}^{\infty} s^y P(Y=y)$$

$$= \sum_{y=0}^{\infty} s^y \left(\frac{1}{2}\right)^{y+1}$$

$$= \frac{1}{2} \cdot \sum_{y=0}^{\infty} \left(\frac{1}{2}s\right)^y$$

$$= \frac{1}{2} \cdot \frac{1}{1 - \frac{1}{2}s} \quad \text{for } \left|\frac{1}{2}s\right| < 1, \text{ i.e. } -2 < s < 2$$

$$G(s) = \frac{1}{2-s} \quad \text{as stated.}$$

This is valid for  $-2 < s < 2$ , so radius of convergence = 2.

b)  $\gamma$  is the smallest solution  $\geq 0$  to  $G(s) = s$ :

$$G(s) = s \Rightarrow \frac{1}{2-s} = s \Rightarrow s^2 - 2s + 1 = 0$$

$$\Rightarrow (s-1)^2 = 0$$

$$\Rightarrow \underline{s = 1.}$$

Thus  $\gamma = P(\text{extinction}) = 1.$

[Alternative: calculate  $\mu = 1 \Rightarrow \gamma = 1.$ ]

c) i)  $P(Z_n = 0) = G_n(0)$

$$= \frac{n - (n-1) \cdot 0}{n+1 - n \cdot 0}$$

$$= \underline{\underline{\frac{n}{n+1}}}$$

- formula correct for  $r=0.$

ii)  $P(Z_n = 1) = G'_n(0)$

Now  $G'_n(s) = (n - (n-1)s)(n+1 - ns)^{-1}$

$$\Rightarrow G'_n(s) = - (n-1)(n+1 - ns)^{-1} - (n - (n-1)s)(n+1 - ns)^{-2} (-n)$$

8

6c cont.) So  $G'_n(0) = -\frac{(n-1)}{n+1} + \frac{n^2}{(n+1)^2}$

$$= \frac{-(n+1)(n-1)}{(n+1)^2} + \frac{n^2}{(n+1)^2}$$

$$= \frac{-n^2 + 1 + n^2}{(n+1)^2}$$

$$= \frac{1}{(n+1)^2}$$

$$= \frac{n^0}{(n+1)^2}$$

- formula correct for  $r=1.$

d)  $P(\text{extinct at generation } 8) = P(Z_8 = 0) - P(Z_7 = 0)$

$$= \frac{8}{9} - \frac{7}{8}$$

$$= \underline{\underline{\frac{1}{72} \cdot (0.0139)}}.$$

e)  $P(Z_{10} \leq 20) = P(Z_{10} = 0) + \sum_{r=1}^{20} P(Z_{10} = r)$

$$= \frac{10}{11} + \sum_{r=1}^{20} \frac{10^{r-1}}{11^{r+1}}$$

using formulas given

$$= \frac{10}{11} + \sum_{s=0}^{19} \frac{10^s}{11^{s+2}}$$

putting  $s = r-1$

$$= \frac{10}{11} + \frac{1}{11^2} \sum_{s=0}^{19} \left(\frac{10}{11}\right)^s$$

- finite sum of Geometric series (in Attachment)

$$= \frac{10}{11} + \frac{1}{121} \left( \frac{1 - \left(\frac{10}{11}\right)^{20}}{1 - \frac{10}{11}} \right)$$

$$\underline{\underline{P(Z_{10} \leq 20) = 0.986}}$$

10

8) a) If  $X_0 \sim (0.1, 0.9)^T$ , then

$$X_1 \sim (0.1, 0.9) \begin{pmatrix} 1-\alpha & \alpha \\ \beta & 1-\beta \end{pmatrix} = \begin{pmatrix} 0.1(1-\alpha) + 0.9\beta \\ 0.1\alpha + 0.9(1-\beta) \end{pmatrix}$$

So  $X_1 \sim (0.1(1-\alpha) + 0.9\beta, 0.1\alpha + 0.9(1-\beta))$

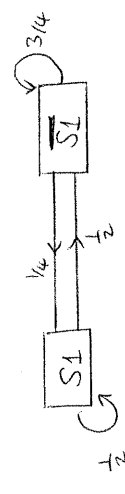
b)  $P(X_1=2, X_2=1, X_3=2 | X_0=2) = p_{22} p_{21} p_{12} = \underline{\underline{(1-\beta)\beta\alpha}}$

c)  $Y_t$  can be in one of two states: State 1 or not state 1. Call these  $S_1$  and  $\bar{S}_1$ .

Now  $P(Y_{t+1}=1 | Y_t=1) = \frac{1}{2}$  from matrix given  $(Q_{11})$ .

Also,  $\begin{cases} P(Y_{t+1}=1 | Y_t=2) = \frac{1}{4} = Q_{21} \\ P(Y_{t+1}=1 | Y_t=3) = \frac{1}{4} = Q_{31} \end{cases}$

So whenever  $Y_t$  is in state  $\bar{S}_1$  (i.e. state 2 or 3),  $P(Y_{t+1}=S_1 | Y_t \in \bar{S}_1) = \frac{1}{4}$ .

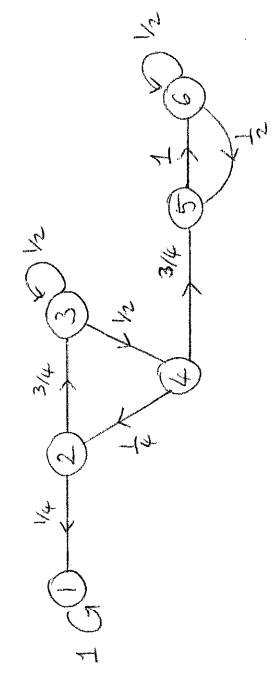


The 2-state matrix is therefore  $P = \begin{pmatrix} \frac{1}{2} & \frac{1}{4} \\ \frac{1}{2} & \frac{3}{4} \end{pmatrix}$  i.e.  $\alpha = \frac{1}{2}$ ,  $\beta = \frac{1}{4}$

So  $P(Y_t=1 | Y_0=1) = P(Y_t=S_1 | Y_0=S_1) = (P^t)_{11} = \frac{1}{\alpha+\beta} \{ \beta + \alpha(1-\alpha-\beta)^t \}$

$= \frac{1}{\frac{1}{2} + \frac{1}{4}} \{ \frac{1}{4} + \frac{1}{2} (\frac{1}{4})^t \}$   
 $= \underline{\underline{\frac{1}{3} + \frac{2}{3} (\frac{1}{4})^t}}$  general solution

9



Communicating classes:  $\{1\}$  closed,  $\{2, 3, 4\}$  not closed,  $\{5, 6\}$  closed.

b)  $h_{iA}$  satisfies:  $h_{iA} = 1$  because  $A = \{1\}$  and  $h_{iA} = \sum_{j=1}^6 p_{ij} h_{jA}$  for  $i \neq 1$ .

Now by inspection,  $h_{5A} = h_{6A} = 0$

Thus  $\begin{cases} h_{2A} = \frac{1}{4} h_{1A} + \frac{3}{4} h_{3A} \\ h_{2A} = \frac{1}{4} + \frac{3}{4} h_{3A} \end{cases} \quad (1)$

and  $h_{3A} = \frac{1}{2} h_{3A} + \frac{1}{2} h_{4A} \Rightarrow h_{3A} = h_{4A} \quad (2)$

and  $h_{4A} = \frac{1}{4} h_{2A} + \frac{3}{4} h_{5A} = \frac{1}{4} h_{2A} \quad (3)$

$\Rightarrow h_{4A} = \frac{1}{4} h_{2A}$

using (3)  $\Rightarrow h_{2A} = \frac{1}{4} + \frac{3}{4} (\frac{1}{4} h_{2A})$

$\frac{13}{16} h_{2A} = \frac{1}{4} \Rightarrow h_{2A} = \frac{4}{13}$

So also,  $h_{3A} = h_{4A} = \frac{1}{4} h_{2A} = \frac{1}{13}$

Thus  $h_{iA} = (1, \frac{4}{13}, \frac{1}{13}, \frac{1}{13}, 0, 0)$

9) The transition matrix has components  $P(Z_{t+1}=j | Z_t=i)$ .

Suppose  $Z_t=0$ : then  $P(Z_{t+1}=0 | Z_t=0) = 1$  (given in Q).

Suppose  $Z_t=1$ : then 1 individual has  $\begin{cases} 0 \text{ offspring (prob } 0) \\ 1 \text{ " (prob } 1) \end{cases}$ .

Suppose  $Z_t=2$ : 2 indivs each have  $\begin{cases} 0 \text{ young (prob. } 1-\frac{1}{2}=\frac{1}{2}) \\ 1 \text{ young (prob. } \frac{1}{2}) \end{cases}$   
 $\Rightarrow Z_{t+1} = \begin{cases} 0 \text{ prob. } (\frac{1}{2})^2 = \frac{1}{4} \\ 1 \text{ prob. } 2 \cdot \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{2} \\ 2 \text{ prob. } (\frac{1}{2})^2 = \frac{1}{4} \end{cases}$

Suppose  $Z_t=3$ : 3 indivs each have  $\begin{cases} 0 \text{ young (prob } 1-\frac{1}{3}=\frac{2}{3}) \\ 1 \text{ young (prob } \frac{1}{3}) \end{cases}$ . Binomial  $(3, \frac{1}{3})$   
 $\Rightarrow Z_{t+1} = \begin{cases} 0 \text{ prob } (\frac{2}{3})^3 = \frac{8}{27} \\ 1 \text{ prob } \binom{3}{1} (\frac{2}{3})^2 (\frac{1}{3}) = \frac{12}{27} \\ 2 \text{ prob } \binom{3}{2} (\frac{2}{3}) (\frac{1}{3})^2 = \frac{6}{27} \\ 3 \text{ prob } (\frac{1}{3})^3 = \frac{1}{27} \end{cases}$

The transition matrix is:  $\begin{matrix} & 0 & 1 & 2 & 3 \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \end{matrix} & \begin{pmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ \frac{1}{4} & \frac{1}{2} & \frac{1}{4} & 0 \\ \frac{8}{27} & \frac{12}{27} & \frac{6}{27} & \frac{1}{27} \end{pmatrix} \end{matrix}$

$P(\text{extinction}) = h_{30} = P(\text{hit state } 0 \text{ starting at state } 3)$ .

Now the hitting probabilities  $\{h_{i0}\}$  satisfy:

$h_{00} = 1$   
 $h_{10} = 0$  by inspection  
 $h_{20} = \frac{1}{4} h_{00} + \frac{1}{2} h_{10} + \frac{1}{4} h_{20}$   
 $\Rightarrow \frac{3}{4} h_{20} = \frac{1}{4} + 0 \Rightarrow h_{20} = \frac{1}{3}$

and  $h_{30} = \frac{8}{27} h_{00} + \frac{12}{27} h_{10} + \frac{6}{27} h_{20} + \frac{1}{27} h_{30}$

$\Rightarrow \frac{26}{27} h_{30} = \frac{8}{27} + 0 + \frac{6}{27} \cdot \frac{1}{3} = \frac{10}{27}$

$P(\text{extinction}) = h_{30} = \frac{5}{13} = \underline{\underline{0.385}}$

10)  $G_{1k}(s) = E(s^{T_k})$ .

Now  $T_1 = 0$  with probability 1, so  $G_1(s) = E(s^{T_1}) = 1$ .

Also  $T_2 = \begin{cases} 1 \text{ with prob. } \frac{1}{4} \\ 1+T_3 \text{ with prob. } \frac{3}{4} \end{cases}$

So  $G_2(s) = E(s^{T_2}) = \frac{1}{4} E(s^1) + \frac{3}{4} E(s^{1+T_3})$

$\therefore G_2(s) = \frac{1}{4} s (1 + 3G_3(s))$  \*

Also  $T_3 = \begin{cases} 1+T_2 \text{ with prob. } \frac{1}{4} \\ 1+T_3 \text{ with prob. } \frac{3}{4} \end{cases}$

So  $G_3(s) = E(s^{T_3}) = \frac{1}{4} E(s^{1+T_2}) + \frac{3}{4} E(s^{1+T_3})$

$= \frac{1}{4} s (G_2(s) + 3G_3(s))$

$= \frac{1}{4} s \left\{ \frac{1}{4} s (1 + 3G_3(s)) + 3G_3(s) \right\}$

$= \frac{1}{16} s^2 + \frac{3}{16} s^2 G_3(s) + \frac{3}{4} s G_3(s)$  from \*

$G_3(s) \left\{ 1 - \frac{3}{16} s^2 - \frac{3}{4} s \right\} = \frac{1}{16} s^2$

$G_3(s) = \frac{s^2}{16 \left( 1 - \frac{3}{16} s^2 - \frac{3}{4} s \right)}$

$G_3(s) = \frac{s^2}{16 - 12s - 3s^2}$